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Aggregated spatial distributions of species in a subtropical karst forest, southwestern China

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Abstract

Aims

Spatial distribution patterns of species reflect not only the ecological processes but also the habitat features that are related to species distribution. In karst topography, species distribution patterns provide more specific information about their environments. The objectives of this study are as follows: (i) to analyse and explain the spatial distribution patterns of conspecific trees in an oldgrowth subtropical karst forest; (ii) to investigate pattern changes at different spatial scales; (iii) to test the spatial pattern similarity (or dissimilarity) between trees at different abundances, diameter at breast height classes, canopy layers and different functional groups (shade tolerance and seed dispersal mode); (iv) to examine whether habitat heterogeneity has an important effect on the species spatial distribution.

Methods

The spatial distributions of woody species with ≥ 20 individuals in a 1-ha subtropical karst forest plot at Maolan in southwestern China were quantified using the relative neighbourhood density Ω based on the average density of conspecific species in a circular neighbourhood around each species.

Important Findings

Aggregated distribution is the dominant pattern in the karst forest, but the ratio of aggregated species in total species number decreases with an increase in spatial scale. Less abundant species are more aggregated than most abundant species. Aggregation is weaker in larger diameter classes, which is consistent with the prediction of self-thinning. Seed dispersal mode influences spatial patterns, with species dispersed by animals being less aggregated than those dispersed by wind and gravity. Other species functional traits (e.g. shade tolerance) also influence the species spatial distributions. Moreover, differences among species habitat associations, e.g. with rocky outcrops, play a significant role in species spatial distributions. These results indicate that habitat heterogeneity, seed dispersal limitation and self-thinning primarily contribute to the species spatial distributions in this subtropical karst forest.

Keywords: functional trait • habitat heterogeneity • karst forest • relative neighbourhood density • spatial distribution
species coexistence.

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INTRODUCTION

The species spatial distribution in forest stands is of particular interest in ecological studies because it reveals information about the stand history, population dynamics and species competition (Haase 1995). Furthermore, it infers the underlying ecological processes or mechanisms (Manabe *et al.*

2000; Nishimura *et al.* 2008; Wiegand *et al.* 2007). Species often distribute aggregately in natural forests (Condit *et al.* 2000; Hardy and Sonké 2004; He *et al.* 1997). In recent decades, numerous mechanisms were proposed to explain the formation of species aggregation in forests. For example, many ecologists successfully demonstrated that plant species are restricted to a more or less wide range of habitats (Itoh

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et al. 2003; Takyu and Kitayama 2002). Large forest plots worldwide show that tree species are closely associated with habitat factors (e.g. topographic and edaphic) and microhabitat heterogeneity created by such factors at local scales in the tropics (e.g. Gunatilleke et al. 2006; Lai et al. 2009; Queenborough et al. 2007; Sri-Ngernyuang et al. 2003; Yamada et al. 2006). Habitat heterogeneity is an important factor in the regulation of species distributions (Harms et al. 2001; Yamada et al. 2006). Species-specific differences in functional traits and ecological strategies also affect the spatial distribution and dynamics of species (Condit et al. 2006; Lortie et al. 2004; Murrell 2009). For example, the seed dispersal ability affects species distributions, with poorly dispersed species being more aggregated than well-dispersed ones (Condit et al. 2000; Li et al. 2009). However, ecologists argued that a single dominant mechanism may not drive patterns of tree species distribution, especially in species-rich forests. Several researchers found that species distributions in plant communities are determined by both habitat heterogeneity and dispersal limitation (Lin et al. 2011; Plotkin et al. 2002; Shen et al. 2009; Wiegand et al. 2007). He et al. (1997) suggested that the distribution patterns of tree species in tropical forests are subjected to multiple controlling factors, e.g. topography, spacing effect, density-dependent processes and species rarity. Therefore, multiple controlling mechanisms may work together to organize the structure and to maintain the high diversity of species-rich forests. Identifying the multiple factors affecting species spatial distribution in various forest ecosystems is essential because it provides valuable insights into processes that structure communities and permit species coexistence.

Existing knowledge about species spatial distributions and underlying mechanisms in species-rich communities is insufficient because most of the research are mainly derived from tropical (Condit et al. 2002; Nishimura et al. 2008; Queenborough et al. 2007; Seidler and Plotkin 2006) and subtropical forests worldwide (e.g. Lai et al. 2009; Legendre et al. 2009; Li et al. 2009). Furthermore, forests in other bioclimatic zones and specific forests within tropical and subtropical zones should be considered in testing the applicability of these theories or hypotheses developed from tropical and subtropical forests. The karst forest in the subtropics, which is characterized by a hyper-diverse habitat heterogeneity and a hyper-complex species composition (e.g. in the subtropical southwestern China; Zhang et al. 2010), is an alternative example. The spatial patterns of tree distributions and their underlying formation mechanisms in such few reported forests can improve our understanding of the mechanisms of species diversity (He et al. 1996) and stable species coexistence (Gunatilleke et al. 2006).

On the other hand, the karst forest in China is mainly distributed in the southwest mountainous regions, where the continuous, largest karst landscape in the world is located (Yuan 1991). The representative subtropical mixed evergreen–deciduous broad-leaved forest in southern Guizhou Province has a specific habitat, a complex community structure and a rich biodiversity (Hou and Jiang 2006; Long *et al.* 2005; Zhu and Wei 1993). The spatial patterns of eight dominant tree species in two forest plots (Zhang *et al.* 2010) showed that six species have significant aggregations at the majority of the scales, whereas the other two species show a random distribution pattern at most scales. Habitat heterogeneity is the key factor in determining the spatial distributions of eight tree species, implying the existence of niche theory in speciesrich karst forests. However, the spatial distribution patterns of other woody species are still unclear.

The large plot size containing detailed information about plant distributions is essential to address species spatial patterns and sustainable extraction of tree species, especially in tropical forests (Condit et al. 1994; Condit 1995). However, studies of the species spatial patterns at a relatively small scale (0-100 m) also helped to identify the relative importance of ecological processes in structuring tree-dominated communities (Getzin et al. 2006; Moeur 1993; Salas et al. 2006; Strasberg 1996; Yasuhiro et al. 2004), especially in places where it is not possible to set-up large-sized plots. This article reports on all of the 43 species spatial distributions with ≥ 20 individuals in a 1-ha karst forest plot at Maolan, southern Guizhou Province. The objectives of this study were as follows: (i) to analyse and explain the spatial distribution patterns of conspecific trees in an old-growth subtropical karst forest; (ii) to investigate the pattern changes at different spatial scales; (iii) to test the spatial pattern similarity (or dissimilarity) between trees at different abundances, diameter at breast height (DBH) classes, canopy layers and different functional groups (shade tolerance and seed dispersal mode); (iv) to examine whether habitat heterogeneity has an important effect on the species spatial distributions. This study is expected to provide important insights into the possible mechanisms that maintain the assemblage of species in a karst forest.

MATERIALS AND METHODS

Study site

The study was conducted in a Man and the Biosphere reserve, namely, the Maolan National Natural Reserve (MNNR) (25°09′20″-25°20′50″N, 107°52′10″-108°05′04″E) located in Libo County, southwestern China's Guizhou Province (Fig. 1a). In this ~20 000-ha reserve, the elevation ranges from 430 to 1078.6 m with an average of 800 m, but the relative peak elevations are generally from 100 to 300 m. Topography is characterized by the typical karst fengcong depression (a combination of clustered peaks, cylindrical and cone-shaped towers, with a common rocky base) (Fig. 1b). Carbonate rocks (mainly of pure limestone and dolomite) are usually exposed on the surface. Therefore, soils are thin and discontinuous in the karst terrain. However, the shallow black limestone soil is rich in organic matter, nutrients (N, P and K) and Ca. This region has a subtropical monsoon climate with plateau climate features such as mild temperature and plenty of rainfall



(b) Karst fengcong-depression landscape

Figure 1: location of the 1 ha plot (a) and the typical karst landscape (b) in the Maolan National Natural Reserve, southwestern China.

throughout the year, but with humid air and cloudy sky. The mean temperature ranges from 8.3°C in January to 26.4°C in July, with an annual mean of 15.3°C. The mean annual frost-free period lasts 315 days. The average annual precipitation is 1320.5 mm. The mean annual relative humidity is 83%. The mean sunshine percentage is only ~20%.

Data collection

The karst terrain is too complex and steep (Fig. 1b) to find a relatively flat area for setting up a bigger forest plot of 16-50 ha similar with many permanent forest plots in the world. Thus, a 1-ha (100×100 m) plot in the summer of 2008 was established in a typical old-growth mixed evergreen–deciduous broad-leaved forest of MNNR (Fig. 1a). To our knowledge, there is no large plot bigger than 1 ha in the subtropical karst forest in the world. This plot is situated near the top of a hill ($25^{\circ}18'07''N$, $107^{\circ}57'21''E$), spanning an altitudinal range of 895-938 m. Two slopes facing southeast and northeast are included. The plot is gently flat in the lower and higher parts, rather moderate or steep in the middle, and with numerous outcrop rocks in the mid-lower part. The mean slope is ~30°. The whole-plot slope ranges from 5° to 48°. The soils are slightly acidic to neutral.

The plot was divided into 100 contiguous 10×10 m subplots using the DQL-1 forest compass (Harbin Optical Instrument Factory, China). Each subplot was further divided into four 5×5 m quadrats. All free-standing woody plants with ≥ 1 -cm DBH were investigated, with each quadrat as the basic unit of vegetation survey. The species names, relative geographic coordinates of each individual and DBH were recorded. Meanwhile, the rock-bareness (basement rock was exposed on the ground surface and not covered by soils) percentage within each quadrat was visually estimated. Vertically, the old-growth karst forest is typically stratified into three vegetation layers and species that occupy different canopy levels at maturity. Thus, the species guilds (groups) were easily distinguished as overstory, midstory and understory tree species (Table 1). The shade tolerance and seed dispersal mode were identified according to the Editorial Board for Flora of Guizhou (1982). A total of 3857 individuals belonging to 58 families, 121 genera and 191 species were documented. The basal area was 23.62 m² ha⁻¹. The plot is dominated by Castanopsis carlesii var. spinulosa, Cyclobalanopsis myrsinifolia, Distylium myricoides, Rhododendron latoucheae, Osmanthus fragrans, Platycarya longipes, Engelhardtia roxburghiana, Rapanea neriifolia and Cyclobalanopsis glauca. The spatial patterns of 43 species with \geq 20 individuals (Table 1) were analysed to obtain a sufficiently large sample size. These species comprised 86.4% of the stems in the plot.

Data analysis

Second-order point pattern analyses are the most widely used methods in quantifying stem-mapped tree distributions. These analyses include the Ripley's *K* function (Ripley 1977) and the relative neighbourhood density function Ω (Condit *et al.* 2000; Wiegand and Moloney 2004), which have been commonly applied to detect individual spatial arrangements within communities and to generate hypotheses regarding the underlying processes controlling the observed patterns (Stoyan and Stoyan 1994; Wiegand *et al.* 2007). The Ripley's *K* function and index Ω are closely related. *K* is the cumulative distribution function of distances between pairs of points, whereas Ω is the derivative of the *K* function and is, thus, a probability density function

Table 1.	properties of 43 woody species with ≥20 individuals in a subtropica	l, mixed evergreen-deciduous broad-leaved karst forest in
Maolan,	southern Guizhou Province	

Species	Family	No. of individuals	Canopy layer	Shade tolerance	Dispersal mode
Acer fabri	Aceraceae	57	Midstory	Shade tolerant	Wind
Adina pilulifera	Rubiaceae	30	Understory	Shade tolerant	Wind
Aidia canthioides	Rubiaceae	168	Understory	Shade tolerant	Gravity
Alniphyllum fortunei	Styracaceae	22	Overstory	Light demanding	Wind
Beilschmiedia kweichowensis	Lauraceae	31	Midstory	Shade tolerant	Gravity
Castanopsis carlesii var. spinulosa	Fagaceae	206	Overstory	Shade tolerant	Gravity
Carpinus pubescens	Betulaceae	34	Overstory	Light demanding	Gravity
Camellia furfuracea	Theaceae	39	Understory	Shade tolerant	Wind
Celtis biondii	Ulmaceae	27	Overstory	Light demanding	Animal
Cinnamomum wilsonii	Lauraceae	62	Midstory	Shade tolerant	Gravity
Clausena dunniana	Rutaceae	92	Understory	Shade tolerant	Gravity
Cyclobalanopsis myrsinifolia	Fagaceae	86	Overstory	Midtolerant	Gravity
Cyclobalanopsis glauca	Fagaceae	84	Overstory	Midtolerant	Gravity
Cyclobalanopsis sessilifolia	Fagaceae	27	Overstory	Midtolerant	Gravity
Daphniphyllum oldhami	Daphniphyllaceae	40	Understory	Shade tolerant	Gravity
Distylium myricoides	Hamamelidaceae	201	Midstory	Light demanding	Wind
Engelhardtia roxburghiana	Juglandaceae	221	Midstory	Light demanding	Wind
Eriobotrya henryi	Rosaceae	34	Midstory	Midtolerant	Wind
Euonymus myrianthus	Celastraceae	60	Understory	Shade tolerant	Gravity
Gardenia jasminoides	Rubiaceae	33	Understory	Shade tolerant	Wind
Ilex ficoidea	Aquifoliaceae	39	Midstory	Shade tolerant	Gravity
Lasianthus japonicus var. lancilimbus	Rubiaceae	97	Understory	Shade tolerant	Gravity
Lindera communis	Lauraceae	157	Understory	Shade tolerant	Wind
Lindera nacusua	Lauraceae	62	Understory	Shade tolerant	Wind
Lindera pulcherrima var. hemsleyana	Lauraceae	33	Understory	Shade tolerant	Animal
Lithocarpus henryi	Fagaceae	61	Overstory	Light demanding	Gravity
Litsea subcoriacea	Lauraceae	24	Midstory	Light demanding	Gravity
Machilus microcarpa	Lauraceae	26	Midstory	Shade tolerant	Animal
Meliosma fordii	Sabiaceae	36	Midstory	Midtolerant	Gravity
Neolitsea aurata	Lauraceae	52	Midstory	Shade tolerant	Animal
Osmanthus fragrans	Oleaceae	183	Midstory	Light demanding	Gravity
Pittosporum crispulum	Pittosporaceae	36	Understory	Shade tolerant	Wind
Pittosporum podocarpum	Pittosporaceae	28	Understory	Shade tolerant	Wind
Platycarya longipes	Juglandaceae	74	Overstory	Light demanding	Wind
Rapanea neriifolia	Myrsinaceae	205	Understory	Shade tolerant	Gravity
Reevesia glaucophylla	Sterculiaceae	73	Midstory	Light demanding	Wind
Rhododendron latoucheae	Ericaceae	174	Midstory	Shade tolerant	Wind
Sloanea sinensis	Elaeocarpaceae	97	Midstory	Midtolerant	Wind
Symplocos lancifolia	Symplocaceae	85	Understory	Shade tolerant	Gravity
Symplocos sumuntia	Symplocaceae	29	Midstory	Shade tolerant	Gravity
Swida austrosinensis	Cornaceae	20	Midstory	Light demanding	Gravity
Tarenna mollissima	Rubiaceae	22	Understory	Light demanding	Gravity
Vaccinium mandarinorum	Ericaceae	21	Understory	Shade tolerant	Animal

(Diggle 2003; Stoyan and Penttinen 2000). The *K* function confounds the effects of large scales with small scales because it is a cumulative measurement and each larger scale includes information from all smaller scales (George *et al.* 2006; Schurr *et al.* 2004; Wiegand *et al.* 2007). Ω isolates specific distance classes

by replacing the Ripley's *K* function circles with rings and using the mean number of neighbours in a ring of radius and width around an individual (Wiegand and Moloney 2004). Thus, the index Ω has an intuitive biological interpretation as a local neighbourhood density (Condit *et al.* 2000). Therefore, the relative neighbourhood density Ω (Condit *et al.* 2000) was used in this study to characterize the tree distributions in the plot. For a given species, Ω is equivalent to the average density of conspecifics in the neighbourhood of individuals, which is normalized by the density of the individual species in the entire plot:

$$\Omega_{\chi 1,\chi 2} = D_{\chi 1,\chi 2}/\lambda$$

where $D_{\chi_1 \prime \chi_2} = \sum N_{\chi_{1,\chi_2}} / \sum A_{\chi_{1,\chi_2}}$. $N_{\chi_1 \prime \chi_2}$ is the number of conspecifics located between distances χ^1 and χ^2 from each individual, averaged by the overall individuals of the species, $A_{\chi_1 \prime \chi_2}$ is the area of the annulus, and λ is the mean density of a given species in the entire plot. For a random distribution, $\Omega_{\chi_1 \prime \chi_2} = 1$ within an annulus between χ^1 and χ^2 meters. $\Omega_{\chi_1 \prime \chi_2} > 1$ indicates aggregation, whereas $\Omega_{\chi_1 \prime \chi_2} < 1$ suggests regular distribution. Complete spatial randomness null models were chosen for analyses. The 99% statistical confidence limit was estimated by performing Monte Carlo procedure with 499 replicates. These calculations were analysed using software SpPack 1.38 (Perry 2004).

 Ω_{0-10} , which refers to the mean conspecific density within 10 m of a tree, was used as the intensity of species aggregation measurement (Condit et al. 2000) to compare species spatial patterns belonging to different characteristic groups. DBH was first used to divide species into six size classes (0-5 cm, 5-10 cm, 10-20 cm, 20-30 cm, 30-40 cm and 40-50 cm). Second, species were classified into overstory, midstory and understory species. The third comparison was performed between species of different seed dispersal modes: wind, gravity and animaldispersed species. Data were tested for normal distributions. The non-normal distributed data were log-transformed to achieve normalization. The differences between groups were then compared using the Student's t-test. A multiple regression for these 43 species of abundance ≥ 20 was conducted to analyse the effects of abundance, DBH, canopy layer, shade tolerance and dispersal mode on spatial patterns using Ω_{0-10} as the dependent variable. Abundance, maximum DBH, average DBH, canopy layer, shade tolerance and dispersal mode were used as the independent variables.

RESULTS

Most of the 43 species in the plot were aggregated at a scale of <50 m (Table 2). Although aggregation was a dominant pattern when all the DBH classes \geq 1 cm were included in the analysis, the percentage of aggregated species decreased with the increase in spatial scale. For example, 93.0% of the species were significantly aggregated from 0 to 10 m, 81.3% of the species from 10 to 20 m and 69.8% of the species from 20 to 30 m (Table 2). The value of Ω invariably decreased with the spatial scale, and the Ω of the less abundant species decreased faster than that of the most abundant species at the smaller scales (<10 m) (Fig. 2).

The aggregation intensity, as measured by Ω_{0-10} , clearly decreased with the abundance in the plot (Fig. 3). For example, the less abundant species with higher and more

Table 2. number of woody species for each spatial distribution pattern and each spatial scale in the Maolan plot, as measured by Ω

<i>r</i> (m)	Aggregated distribution	Random distribution	Regular distribution
0-10	40	3	0
10–20	35	8	0
20-30	30	13	0
30-40	26	13	4
40-50	24	14	5

dispersed Ω_{0-10} values were more aggregated than the most abundant species having $\Omega_{0-10} < 5$. The largest Ω_{0-10} was 31.1 (*Alniphyllum fortunei*, 22 individuals; Fig. 2a).

The aggregation intensity varied for the different DBH size classes. The smaller size classes showed a higher degree of aggregation than the larger size classes. The percentage of significantly aggregated species decreased with the DBH (Table 3). Specifically, 95.3% of the species at the 1–5 cm DBH size class, 75% at the 20–30 cm DBH size class and 33.3% at the 40–50 cm DBH size class were aggregated. Taking the two dominant species *C. carlesii* var. *spinulosa* and *C. glauca* as examples, the aggregation intensity of these species declined with the DBH (Fig. 4), indicating that the larger trees were more dispersed than the smaller trees.

The average Ω_{0-10} of the overstory species (5.8, SE = 1.6) was lower than those of the midstory species (6.9, SE = 0.9) and the understory species (9.1, SE = 1.8). The results of the *t*-test only showed a significant difference between the overstory and understory species. The overstory species were less aggregated than the understory species, whereas the midstory species were not significantly different from both the overstory and the understory species.

The average Ω_{0-10} of the light-demanding species (6.3, SE = 1.3) was lower than that of shade-tolerant ones (8.2, SE = 1.2), whereas the mid-tolerant species had the lowest average Ω_{0-10} (4.3, SE = 0.8). The *t*-test showed that the aggregation intensity of the mid-tolerant species was significantly different from those of the shade-tolerant and light-demanding species. The mid-tolerant species were less aggregated than the shade-tolerant and light-demanding species, whereas the light-demanding species were not significantly different from the shade-tolerant species.

The average Ω_{0-10} of the animal-borne species (4.4, SE = 0.7) was lower than that of the gravity-borne species (6.6, SE = 0.9) and the wind-borne species (8.9, SE = 1.8). The *t*-test showed that only the aggregation intensity of the animal-dispersed species was significantly different from those of the gravity-dispersed and wind-dispersed species. The species that are dispersed by animals were less aggregated than those dispersed by gravity or the wind. In contrast, the gravity-dispersed species were not significantly different from the wind-dispersed species.

The multiple regression analysis of Ω_{0-10} showed that the regression model was highly significant (analysis of



Figure 2: examples of species distribution in the Maolan plot. Left column shows the relationship between Ω and the spatial scale for six species; right column shows their corresponding distribution patterns. The six species were chosen from high to low abundance. The black line represents the Ω value. The thin dashed line corresponds to the confidence intervals generated from 499 Monte Carlo simulations under the null hypothesis of complete spatial randomness.



Figure 3: relationship between abundance and Ω_{0-10} of woody species with abundance ≥ 20 at the Maolan plot with the following DBH classes: (1) 1–5 cm, (2) 5–10 cm, (3) 10–20 cm, (4) 20–30 cm, (5) 30–40 cm, and (6) 40–50 cm.

variance, *F*-test with *P*-value = 0.011). The standardized coefficients indicated that the abundance had the largest effect on the spatial aggregation, followed by the maximum DBH, the average DBH, the dispersal mode, the level of shade tolerance and the canopy layer (Table 4). Except for the average DBH and the dispersal mode, the effects of all factors on aggregation were negative, i.e. the aggregation intensity decreased with an increase in the respective factors.

Table 3. number of woody species in different DBH classes with

 significantly aggregated spatial distribution in the Maolan plot

DBH class (cm)	Median Ω_{0-10}	Total no. of species	No. of significant aggregated species
1–5	9.35	43	41
5-10	8.17	42	38
10-20	6.58	36	31
20-30	4.55	20	15
30-40	2.14	7	4
40-50	1.02	3	1



Figure 4: relationship between Ω_{0-10} and the DBH of *C. carlesii* var. *spinulosa* and *C. glauca* with the following DBH classes: (1) 1–5 cm, (2) 5–10 cm, (3) 10–20 cm, (4) 20–30 cm, (5) 30–40 cm, and (6) 40–50 cm.

Table 4. multiple regression of Ω_{0-10} with the abundance, maximum DBH, average DBH, canopy layer, shade tolerance, and dispersal mode showing the estimated coefficients, standard errors and standardized coefficients

	Unstandardized coefficients		Standardized	
	Estimates	SE	(beta) coefficients	
Constant	8.661	5.174		
Abundance	-0.041	0.015	-0.437	
Max DBH	-0.178	0.132	-0.374	
Average DBH	0.582	0.435	0.346	
Canopy layer	-0.130	1.697	-0.017	
Shade tolerance	-0.406	1.105	-0.063	
Dispersal mode	1.513	1.209	0.173	

The standardized coefficients are partial regression coefficients that indicate the relative effects of each variable on Ω_{0-10} .

DISCUSSION

Any population in a community, at a given scale of observation, presents one of three distributions, aggregated, random or regular, depending on the underlying processes (He et al. 1997). Few species in nature have a regular distribution. On the contrary, aggregation is a common pattern of species distribution, particularly in species-rich tropical forests (Itoh et al. 1997; Plotkin et al. 2000). The species in the subtropical karst forest considered in this study followed this general pattern, but the proportion of aggregated species decreased with the increasing spatial scale. Aggregation occurred in 93.0% of the species at 0-10 m, 81.3% at 10-20 m and 69.8% at 20–30 m (Table 2). Similar results were also found in a temperate forest (Wang et al. 2010), but a clear decrease in species aggregation accompanying the increased spatial scale was not found in tropical and subtropical forests (Condit et al. 2000; Li et al. 2009). In an earlier study, the percent aggregation at all scales were all higher than 97.8% in tropical forests, based on counting the aggregation patterns of 1768 species with at least one individual per hectare (Condit et al. 2000). In subtropical

forests, the aggregation percentages were also all >96.1% at all scales (Li *et al.* 2009).

Numerous factors play important roles in determining the spatial patterns of trees in a plant community. The spatial distribution of tree species can arise from many biotic and abiotic processes such as regeneration, competition, dispersal limitation, habitat heterogeneity, disturbances and other stochastic events (Camarero *et al.* 2000; Hou *et al.* 2004; Houle 1994; Palmiotto *et al.* 2004; Pelissier 1998;). This study showed that the abundance of a species, DBH class, canopy layer, functional traits (shade tolerance and seed dispersal mode) and habitat heterogeneity are important factors affecting the spatial patterns of tree species in a subtropical karst forest.

The degree of aggregation of a species was strongly inversely correlated with its abundance. The clumping intensity decreased as the individual abundance increased (Fig. 3), which is consistent with other forest studies showing that rare species are more clumped than common species (Condit *et al.* 2000). However, not all species have a similar response. In our study, e.g. the abundant species *Clausena dunniana* with 92 individuals had a relatively high Ω_{0-10} of 10.2 (Fig. 2e). One of the most important reasons for this observation is that the spatial distribution of a species can arise from a strong habitat preference (Zhang *et al.* 2010).

The species aggregation generally decreased with increasing DBH in the plot. The observation that aggregation was weaker in classes with larger diameters may be due to the self-thinning process or their density-dependent mortality. Water and soil have been shown to be limiting resources in karst forests (Zhang *et al.* 2010), and adjacent stems in the studied forest are likely to compete with each other for these resources. A number of studies support that there is less aggregation with increasing DBH (Condit *et al.* 2000; Davis *et al.* 2005; He *et al.* 1997). Our result is consistent with the findings from tropical and subtropical forests (Condit *et al.* 2000; Li *et al.* 2009), which report that small-size classes are more clumped than larger-size classes. This trend suggests that the larger trees competitively inhibit conspecifics over a larger area of influence than the smaller trees. Furthermore, the tree spacing patterns are influenced by the tree size and the tree species.

Several studies suggested that canopy trees usually have well-dispersed seeds relative to the understory treelets; these trees are thus assumed to be less clumped than the treelets (Condit *et al.* 2000; Wang *et al.* 2010). A better dispersal-reducing aggregation was found in the studied forest. There was a significant difference in the aggregation intensity between the overstory and the understory species in the subtropical karst forest. The overstory species tend to be less aggregated than the understory species.

The shade tolerance of a species may also be expected to have a significant effect on distribution pattern. Lorimer (1980) and Wang et al. (2009) demonstrated that the midtolerant species have almost unimodal distributions with suppressed small trees, whereas the shade-tolerant species tend to have a steeply descending monotonic diameter distribution with a number of suppressed small trees. This observation indicates that the smaller trees are more aggregated than the larger trees. Thus, the shade-tolerant species are expected to be more aggregated than the mid-tolerant species. Lightdemanding species also have the tendency to live in gaps created by tree-falls, thereby causing more aggregation than the mid-tolerant species. In this study, the mid-tolerant species were also found to be less aggregated than the shade-tolerant species and the light-demanding species in the subtropical karst forest.

Several studies found that dispersal limitation is commonly regarded as one of the important mechanisms to explain species aggregation (Plotkin et al. 2000; Seidler and Plotkin 2006) in tropical forests (Condit et al. 2000; He et al. 1997), subtropical forest (Li et al. 2009) and temperate forests (Hou et al. 2004; Wang et al. 2010). Several studies showed that the extent and scale of conspecific spatial aggregation is correlated with the mode of seed dispersal. Animal-dispersed species generally have a high dispersal ability compared with either the wind- or gravity-dispersed species; thus, the species dispersed by animals are assumed to be less clumped than species dispersed by gravity or the wind (Condit et al. 2000; Li et al. 2009). This study also showed that the animalborne species are better dispersed than both the wind- and gravity-borne species. Furthermore, the animal-dispersed species were less aggregated than the wind- and gravity-dispersed species. Thus, the clustering of these species can be linked to their dispersal properties, which are associated with the random interception and accumulation of seeds by the nearby neighbouring vegetation, including the trees, shrubs and dense understory grasses. The aggregated distributions of numerous species indicate the existence of a dispersal limitation in the subtropical karst forest. However, the karst forest has a rocky outcrop and a steep slope. Therefore, some of the species occurring in small-scale clumps do correspond with the topographic and edaphic factors.

Habitat heterogeneity, which involves topographic, edaphic or other environmental factors, is considered an important factor that controls the distribution of a species (Harms *et al.*



Figure 5: Spatial distribution of *O. fragrans, D. myricoides, L. japonicus* var. *lancilimbus* and *C. dunniana* in different habitats. The rock-bareness rate in each 5×5 m quadrat is shown in different grey levels throughout the entire plot.

2001; Yamada et al. 2006). In the study site of this research, the topography such as rocky outcrops, steep slopes, elevation and slope aspect lead to the high habitat heterogeneity of the karst forest. Most tree species in the plot showed significant topographic habitat associations. The prevalence of rock is associated with a low water-holding capacity and shallow soil. Trees on rocky slopes are susceptible to water shortage (Zhou and Pan 2001). Shallow soils interspersed with rocky outcrops limit the availability of nutrients for tree growth (Huang et al. 2009). These conditions may lead to habitat specialization. Thus, species in the plot were aggregated due to habitat specificity, and each species had its preferred habitat. For example, O. fragrans and Lasianthus japonicus var. lancilimbus favour habitats with relatively less rocky outcrops and shady slopes. However, D. myricoides and C. dunniana are established more successfully in habitats with rocky outcrops and sunny slopes (Fig. 5). Habitat specialization that is based on the niche differentiation of resources can explain why different tree species are best suited to different habitats (Harms et al. 2001). Therefore, coexisting species are strongly associated with and are adapted to various habitats that are required for their survival and recruitment. Such habitat heterogeneity is another key reason for the spatial patterns observed in the Maolan plot.

CONCLUSIONS

In summary, the results of this study show that aggregation is dominant at the plot scale (1 ha) for tree species in the speciesrich subtropical karst forest of the Maolan plot. The aggregation intensity clearly declines with the increase in the spatial scale, and the less abundant species are more aggregated than the most abundant species. The aggregation intensity of a species also declines with increasing DBH, which is consistent with the prediction of self-thinning. The functional traits of a species (such as shade tolerance and seed dispersal mode) also have an effect on the spatial distribution of species. In the karst landscape, the topographic habitat partitioning also plays a significant role in the spatial distribution of a species. Therefore, habitat heterogeneity, seed dispersal limitation and self-thinning may have primarily contributed to the observed spatial distribution of various species in the karst forest.

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REFERENCES

- Camarero JJ, Gutierrez E, Fortin MJ (2000) Spatial pattern of subalpine forest-alpine grassland ecotones in the Spanish Central Pyrenees. *For Ecol Manage* **134**:1–16.
- Condit R (1995) Research in large, long-term tropical forest plots. *Trends Ecol Evol* **10**:19–22.
- Condit R, Ashton PS, Baker P, *et al.* (2000) Spatial patterns in the distribution of tropical tree species. *Science* **288**:1414–8.
- Condit R, Ashton P, Bunyavejchewin S, *et al.* (2006) The importance of demographic niches to tree diversity. *Science* **313**:98–101.
- Condit R, Hubbell SP, Foster RB (1994) Density dependence in two understory tree species in a neotropical forest. *Ecology* **75**:671–80.
- Condit R, Pitman N, Leigh EG Jr, *et al.* (2002) Beta-diversity in tropical forest trees. *Science* **295**:666–9.
- Davis MA, Curran C, Tietmeyer A, *et al.* (2005) Dynamic tree aggregation patterns in a species-poor temperate woodland disturbed by fire. *J Veg Sci* **16**:167–74.
- Diggle PJ (2003) Statistical Analysis of Spatial Point Patterns (2nd ed). London, UK: Arnold.
- Editorial Board for Flora of Guizhou. (1982) *Flora of Guizhou*. Guiyang, China: Guizhou People Press.
- George LWP, Miller BP, Enright NJ (2006) A comparison of methods for the statistical analysis of spatial point patterns in plant ecology. *Plant Ecol* **187**:59–82.
- Getzin S, Dean C, He FL, *et al.* (2006) Spatial patterns and competition of tree species in a Douglas-fir chronosequence on Vancouver Island. *Ecography* **29**:671–82.
- Gunatilleke CVS, Gunatilleke IAUN, Esufali S, *et al.* (2006) Species– habitat associations in a Sri Lankan dipterocarp forest. *J Trop Ecol* **22**:371–84.
- Haase P (1995) Spatial pattern analysis in ecology based on Ripley's K-function: introduction and methods of edge correction. *J Veg Sci* 6:575–82.
- Hardy OJ, Sonké B (2004) Spatial pattern analysis of tree species distribution in a tropical rain forest of Cameroon: assessing the role of limited dispersal and niche differentiation. *For Ecol Manag* **197**:191–202.
- Harms KE, Condit R, Hubbell SP, et al. (2001) Habitat associations of trees and shrubs in a 50-ha neotropical forest plot. J Ecol 89:947–59.
- He FL, Legendre P, Lafrankie JV (1996) Spatial pattern of diversity in a tropical rain forest in Malaysia. *J Biogeogr* **23**:57–74.
- He FL, Legendre P, Lafrankie JV (1997) Distribution patterns of tree species in a Malaysian tropical rain forest. *J Veg Sci* **8**:105–14.
- Hou JH, Mi XC, Liu CR, *et al.* (2004) Spatial patterns and associations in a *Quercus-Betula* forest in northern China. *J Veg Sci* **15**:407–14.
- Hou MF, Jiang ZC (2006) Species diversity of karst original forest in different geochemical environments in Maolan. *Ecol Env Sci* 15: 572–576 (in Chinese).
- Houle G (1994) Spatiotemporal patterns in the components of regeneration of four sympatric tree species—*Acer rubrum, Asaccharum, Relula alleghaniensis* and *Fagus grandifolia. J Ecol* **82**:39–53.

- Huang YQ, Zhao P, Zhang ZF, *et al.* (2009) Transpiration of *Cyclobalanopsis glauca* (syn. *Quercus glauca*) stand measured by sap-flow method in a karst rocky terrain during dry season. *Ecol Res* **24**:791–801.
- Itoh A, Yamakura T, Ogino K, *et al.* (1997) Spatial distribution patterns of two predominant emergent trees in a tropical rainforest in Sarawak Malaysia. *Plant Ecol* **132**:121–36.
- Itoh A, Yamakura T, Ohkubo T, *et al.* (2003) Importance of topography and soil texture in the spatial distribution of two sympatric dipterocarp trees in a Bornean rainforest. *Ecol Res* **18**:307–20.
- Lai JS, Mi XC, Ren HB, *et al.* (2009) Species-habitat associations change in a subtropical forest of China. *J Veg Sci* **20**:415–23.
- Legendre P, Mi X, Ren H, *et al.* (2009) Partitioning beta diversity in a subtropical broad-leaved forest of China. *Ecology* **90**:663–74.
- Li L, Huang ZL, Ye WH, *et al.* (2009) Spatial distributions of tree species in a subtropical forest of China. *Oikos* **118**:495–502.
- Lin YC, Chang LW, Yang KC, *et al.* (2011) Point patterns of tree distribution determined by habitat heterogeneity and dispersal limitation. *Oecologia* **165**:175–84.
- Long CL, Yu SX, Xing ZB, et al. (2005) Species diversity and regeneration in forest gaps of the Karst forest in Maolan National Nature Reserve, Guizhou Province. *Biodivers Sci* 13:43–50.
- Lorimer CG (1980) Age structure and disturbance history of a southern Appalachian virgin forest. *Ecology* **65**:1169–84.
- Lortie CJ, Brooker RW, Choler P, *et al.* (2004). Rethinking plant community theory. *Oikos* **107**:433–8.
- Manabe T, Nishimura N, Miura M, *et al.* (2000) Population structure and spatial patterns for trees in a temperate old-growth evergreen broad-leaved forest in Japan. *Plant Ecol* **151**:181–97.
- Moeur M (1993) Characterizing spatial patterns of trees using stemmapped data. *Forest Sci* **39**:756–75.
- Murrell DJ (2009). On the emergent spatial structure of size-structured populations: when does self-thing lead to a reduction in clustering? *J Ecol* **97**:256–66.
- Nishimura S, Yoneda T, Fujii S, *et al.* (2008) Spatial patterns and habitat associations of Fagaceae in a hill dipterocarp forest in Ulu Gadut, West Sumatra. *J Trop Ecol* **24**:535–50.
- Palmiotto PA, Davies SJ, Vogt KA, *et al.* (2004) Soil-related habitat specialization in dipterocarp rain forest tree species in Borneo. *J Ecol* **92**:609–23.
- Pelissier R (1998) Tree spatial patterns in three contrasting plots of a southern Indian tropical moist evergreen forest. *J Trop Ecol* **14**:1–16.
- Perry GLW (2004) SpPack: spatial point pattern analysis in Excel using Visual Basic for Applications (VBA). *Environ Model Softw* 19:559–69.
- Plotkin JB, Chave J, Ashton PS. (2002) Cluster analysis of spatial patterns in Malaysian tree species. *Am Nat* **160**:629–44.
- Plotkin JB, Potts MD, Leslie N, *et al.* (2000) Species-area curves, spatial aggregation, and habitat specialization in tropical forests. *J Theor Biol* **207**:81–99.
- Queenborough SA, Burslem DFRP, Garwood NC, *et al.* (2007) Habitat niche partitioning by 16 species of Myristicaceae in Amazonian Ecuador. *Plant Ecol* **192**:193–207.

Ripley BD (1977) Modeling spatial patterns. JR Stat Soc B 39:172–212.

- Salas C, LeMay V, Núñez P, *et al.* (2006) Spatial patterns in an oldgrowth Nothofagus obliqua forest in south-central Chile. *For Ecol Manage* **231**:38–46.
- Schurr FM, Bossdorf O, Milton SJ, *et al.* (2004) Spatial pattern formation in semi-arid shrubland: a priori predicted versus observed pattern characteristics. *Plant Ecol* **173**:271–82.
- Seidler TG, Plotkin JB (2006) Seed dispersal and spatial pattern in tropical trees. *Plos Biol* **4**:2132–8.
- Shen G, Yu M, Hu XS, *et al.* (2009) Species-area relationships explained by the joint effects of dispersal limitation and habitat heterogeneity. *Ecology* **90**:3033–41.
- Sri-Ngernyuang K, Kanzaki M, Mizuno T, et al. (2003) Habitat differentiation of Lauraceae species in a tropical lower montane forest in northern Thailand. Ecol Res 18:1–14.
- Stoyan D, Penttinen A (2000) Recent applications of point process methods in forestry statistics. *Stat Sci* **15**:61–78.
- Stoyan D, Stoyan H (1994) Fractals, Random Shapes and Point Fields: Methods in Geometrical Statistics. Chichester, UK: John Wiley & Sons.
- Strasberg D (1996) Diversity, size composition and spatial aggregation among trees on a 1-ha rain forest plot at La Réunion. *Biodivers Conserv* 5:825–40.
- Takyu MAS, Kitayama K (2002) Effects of topography on tropical lower montane forests under different geological conditions on Mount Kinabalu, Borneo. *Plant Ecol* 159:35–49.
- Wang XG, Hao ZQ, Zhang J, *et al.* (2009) Tree size distributions in an old-growth temperate forest. *Oikos* **118**:25–36.
- Wang XG, Ye J, Li BH, *et al.* (2010) Spatial distributions of species in an old-growth temperate forest, northeastern China. *Can J For Res* **40**:1011–19.
- Wiegand T, Gunatilleke S, Gunatilleke N. (2007) Species associations in a heterogeneous Sri Lankan dipterocarp forest. Am Nat 170:E77–95.
- Wiegand T, Moloney KA (2004) Rings, circles and null-models for point pattern analysis in ecology. *Oikos* **104**:209–29.
- Yamada T, Tomita A, Itoh A, *et al.* (2006) Habitat associations of Sterculiaceae trees in a Bornean rain forest plot. *J Veg Sci* 17:559–66.
- Yasuhiro K, Hirofumi M, Kihachiro K (2004) Effects of topographic heterogeneity on tree species richness and stand dynamics in a subtropical forest in Okinawa Island, southern Japan. J Ecol 92:230–40.
- Yuan DX (1991) Karst of China. Beijing, China: Geological Publishing House, 1–9.
- Zhang ZH, Hu G, Zhu JD, *et al.* (2010) Spatial patterns and interspecific associations of dominant tree species in two old-growth karst forests, SW China. *Ecol Res* **25**:1151–60.
- Zhou YC, Pan GX (2001) Adaptation and adjustment of Maolan forest ecosystem to karst environment. *Carsologica Sin* **20**:47–52.
- Zhu SQ, Wei LM (1993) Phytocoenological characteristics of Maolan karst forest. In Zhu SQ (ed) *Ecological Research on Karst Forest Ecosystem I.* Guiyang: Guizhou Science and Technology Press, 1–11.